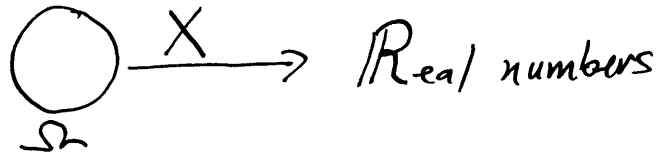


Chapter #3: Discrete random variables and their distributions functions

- A random variable (r.v.) X is a numeric description of an experimental outcome.

More sophisticatedly said, a r.v. X is a real-valued function defined on a sample space Ω .



- The range of X is the set of values $\{X(\omega) : \omega \in \Omega\}$

and we denote this by R_X . $R_X \subseteq \mathbb{R}$.

The r.v. X is a discrete r.v. if R_X is discrete; i.e., $R_X = \{x_1, x_2, \dots, x_n, \dots\}$. Hence, its graph on the real no. line looks like



points with well defined gaps between them.

The event $\{X=x\}$ has a probability mass associated with it; it is denoted by

$$P(X=x).$$

Of course, if $x \notin R_X$, then $\{X=x\} = \emptyset$ and

$$P(X=x) = P(\emptyset) = 0.$$

- The probability function (pf) of the r.v. X , denoted by $f_x(x)$, is defined as

$$f_x(x) = \begin{cases} P(X=x) & , \text{ if } x \in R_x \\ 0 & , \text{ if } x \notin R_x. \end{cases}$$

(Some authors use the term probability mass function; pmf.)

Of course, $\sum_{x \in R_x} f_x(x) = 1$

Since $R_x = \{X(\omega) = x : \omega \in \Omega\}$
 = all possible values X can assume.

- The distribution function (df), $F_x(x)$, of a r.v. X is a real-valued function and is defined as

$$F_x(x) = P(X \leq x) \\ = \sum_{\text{all } x_j \leq x} f_x(x_j) = \sum_{\text{all } x_j \leq x} P(X=x_j)$$

since discrete

This is a step-function and steps up at each possible value X may assume. The size of the jump is $P(X=x_j)$.

Note: In general, $f(x)$ denotes a p.f.
and $F(x)$ denotes a d.f.

Example:

① Experiment: Toss a "fair" coin three times.
Let X = number of heads for each outcome.

$$\Omega = \left\{ \begin{array}{cccc} \omega_1 & \omega_3 & \omega_5 & \omega_7 \\ TTT & THT & HHT & TTH \\ \omega_2 & \omega_4 & \omega_6 & \omega_8 \\ TTH & HTT & HTH & HHH \end{array} \right\}$$

$$R_x = \{ X(\omega); \omega \in \Omega \} = \{0, 1, 2, 3\}$$

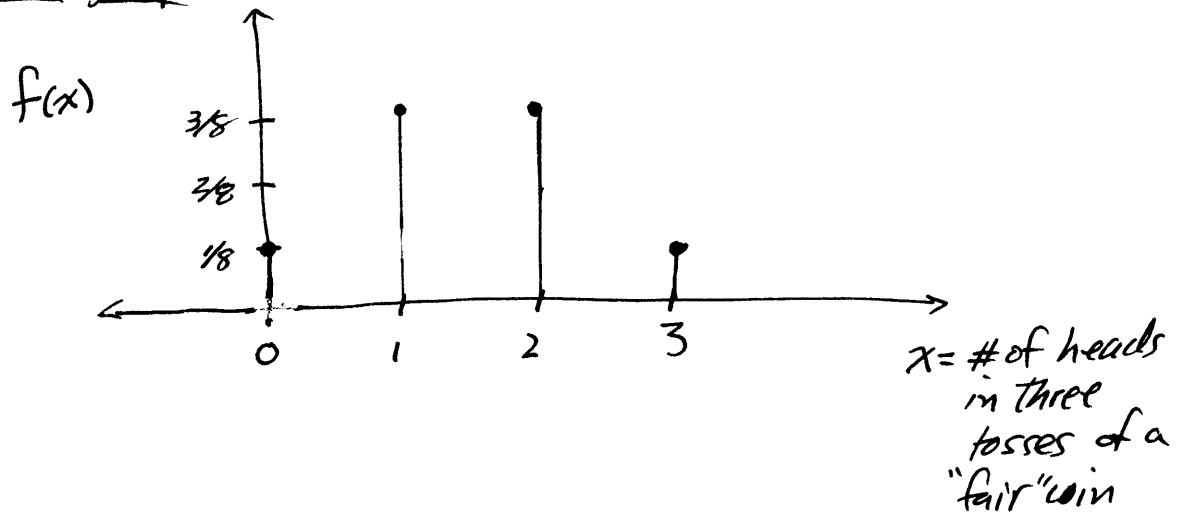
$P(\omega_j \in \Omega) = \frac{1}{8}$
since
equally likely
outcomes

$X=x$	$f(x) = P(X=x)$
0	$\frac{1}{8}$
1	$\frac{3}{8}$
2	$\frac{3}{8}$
3	$\frac{1}{8}$
	$\frac{8}{8} = 1.$

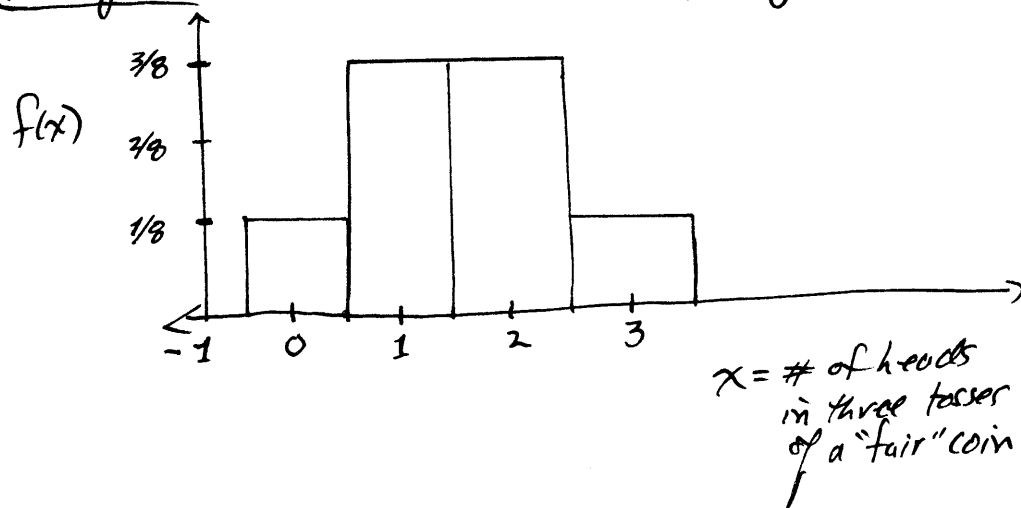
← This is a tabular display of this probability distribution of the r.v. X .

4.

- line graph as a visual display



- Histogram as a visual display



- Mathematical formula to describe a probability distribution

$$f(x) = P(X=x) = \binom{3}{x} \left(\frac{1}{2}\right)^3, \quad x=0, 1, 2, 3$$

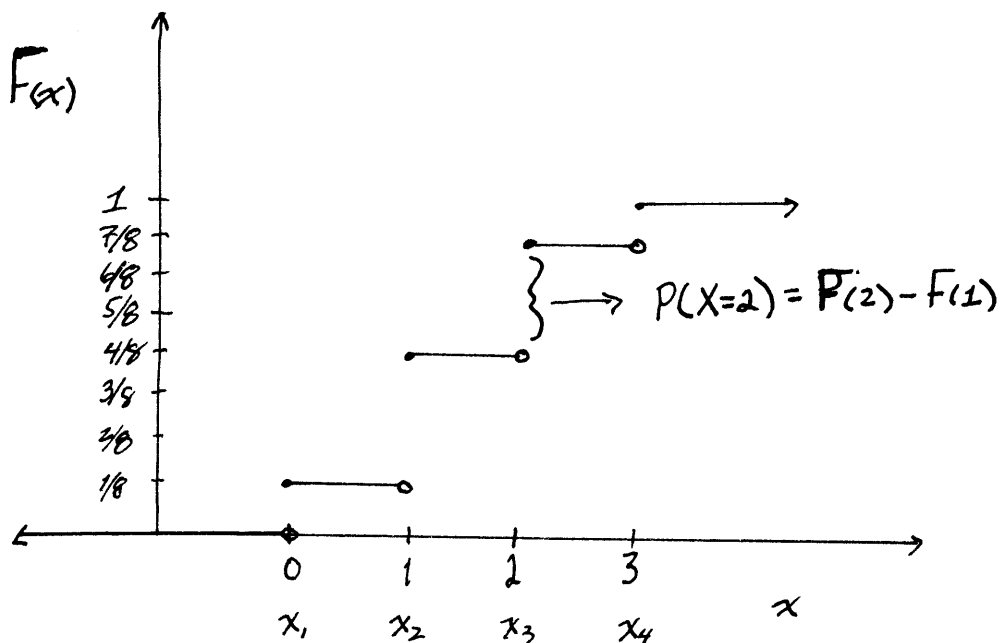
check this: $f(0) = \binom{3}{0} \left(\frac{1}{2}\right)^3 = \frac{3!}{0!3!} \cdot \frac{1}{8} = \frac{1}{8} \checkmark$

$f(1) = \binom{3}{1} \left(\frac{1}{2}\right)^3 = \frac{3!}{1!2!} \cdot \frac{1}{8} = \frac{3}{8} \checkmark$

$f(2) = \binom{3}{2} \left(\frac{1}{2}\right)^3 = \frac{3}{8} \checkmark$

$f(3) = \binom{3}{3} \left(\frac{1}{2}\right)^3 = \frac{1}{8} \checkmark$

• The d.f. $F(x)$ for this example:



$$R_x = \{x_1, x_2, x_3, x_4\} = \{0, 1, 2, 3\}$$

$$F(x) = \begin{cases} 0, & \text{if } x < 0 \\ 1/8, & \text{if } 0 \leq x < 1 \\ 4/8, & \text{if } 1 \leq x < 2 \\ 7/8, & \text{if } 2 \leq x < 3 \\ 1, & \text{if } 3 \leq x < \infty \end{cases}$$

$F(x)$ is increasing, from $0 \rightarrow 1$, step function. It jumps up at each value in R_x and is right continuous

Note that $F(x)$ is defined now over the entire real number line.

Note: $f(x_j) = P(X=x_j) = F(x_j) - F(x_{j-1})$;

of course, we assume the range is written in increasing order so that

$$x_1 < x_2 < \dots < x_{j-1} < x_j < \dots < x_n < \dots$$

$F(x)$ is the r.v. analogue to the ~~own~~ relative cumulative ogive and to the empirical distribution function $\hat{F}_n(x)$.

Note then:

① If X is a discrete r.v. with pf $f_X(x)$, then

$$f_X(x) = P(X=x) \text{ for } x \in R_X.$$

The pf $f_X(x) \geq 0$ and $\sum_{\text{all } x \in R_X} f_X(x) = 1.$

② Let $x_1 < x_2 < \dots < x_{j-1} < x_j < \dots < x_n < \dots$ denote values x in R_X . Then the df $F_X(x)$ determines the values $f_X(x)$ since

$$f_X(x_j) = P(X=x_j) = F_X(x_j) - F_X(x_{j-1}).$$

Let $Q(X)$ denote a function of a discrete r.v. X .
Let X have the following probability distribution:

x	-2	-1	0	1	2
$f(x)$	0.1	0.2	0.3	0.3	0.1

Compute the pf of $Y = X^2$. (Here $y = Q(x) = x^2$.)

y	0	1	4
$f_Y(y)$	0.3	0.5	0.2

Since, e.g., $P(Y=4) = P(X=-2 \text{ or } 2)$
 $= P(X=-2) + P(X=2)$
 $= 0.1 + 0.1 = 0.2$

I.e., the event $Y=4$ occurs when and only when $X=-2$ or $X=2$ occurs.

Recall

the sample mean $\bar{X} = \frac{\sum_{i=1}^n x_i}{n}$

the sample mean for data grouped into a frequency table where x is the distinct value

$f(x)$ = number of times x occurs in the data set

$\hat{f}(x) = \frac{f(x)}{n}$ = the relative frequency.

Then $\bar{X} = \sum x \hat{f}(x) = \frac{\sum x f(x)}{n}$

• Definition:

The expected value (or mean) of a discrete r.v. X with pf $f_x(x)$ is ~~denoted~~ denoted $E(X)$ and is defined to be

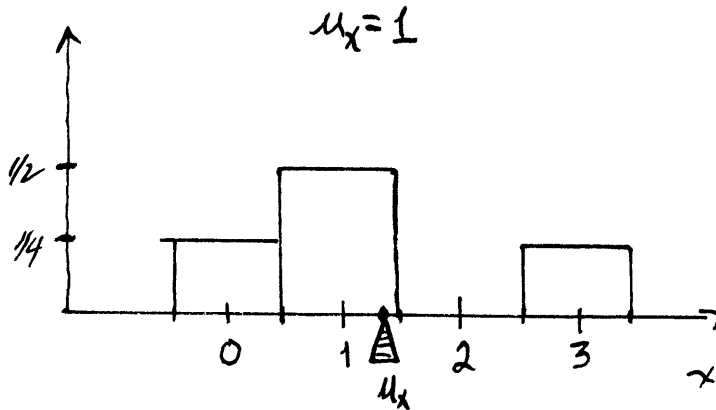
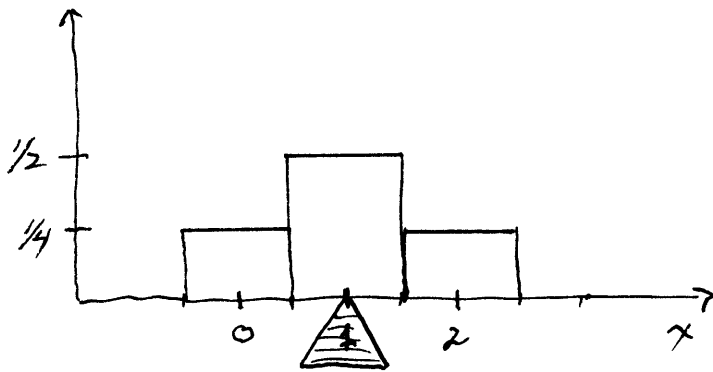
$$E(X) = \sum_{\text{all } x} x P(X=x) = \sum_x x f_x(x).$$

Often we use $\boxed{\mu_x = E(X)}$.

Note: $E(X) < \infty$ provided $\sum_x |x| f_x(x) < \infty$.

If $E(X) < \infty$, we say the mean exists. If $E(X) = \infty$, we say it does not exist.

$E(X)$ is the value where the histogram balances. I.e., It is the center of mass of a distribution of weights $f(x_i)$ located at the points x_i on the line.



Here, $E(X) = \mu_x = 0\left(\frac{1}{4}\right) + 1\left(\frac{1}{2}\right) + 3\left(\frac{1}{4}\right) = \frac{5}{4} = 1.25$

• $E(\varphi(X)) = \sum_{\text{all } x} \varphi(x) f_x(x)$. Says we do not need to know the pf of $Y = \varphi(X)$.

Back to page 6 $\Rightarrow Y = X^2$

$$\begin{aligned}
 E(Y) &= E(X^2) = \sum_x x^2 f_x(x) \\
 &= (-2)^2(.1) + (-1)^2(.2) + 0(.3) + 1^2(.3) + 2^2(.1) \\
 &= .4 + .2 + .3 + .4 = 1.3
 \end{aligned}$$

or $E(Y) = \sum_y y f_y(y) = 0(.3) + 1(.5) + 4(.2)$
 $= 0 + .5 + .8 = 1.3$

- The K^{th} moment μ_K , $K=1, 2, 3, \dots$ of a ~~r.v.~~ r.v. X is defined to be

$$\mu_K = E(X^K) \quad \text{where } K=1, 2, 3, \dots$$

$$= \sum_x x^K f_x(x) \quad \text{when } X \text{ is discrete}$$

Note: Here $Q(x) = x^K$.

- Some facts from intro. calculus or even algebra II:

$$\textcircled{1} \quad \sum_{i=1}^N i = \frac{N(N+1)}{2}$$

$$\textcircled{2} \quad \sum_{i=1}^N i^2 = \frac{N(N+1)(2N+1)}{6}$$

$$\textcircled{3} \quad \sum_{x=0}^{\infty} p^x = 1 + p + p^2 + p^3 + \dots$$

for $0 < p < 1$,

$$= \frac{1}{1-p}$$

This is just the geometric series.

- The variance of a r.v. X is defined to be

$$E[(X - \mu_x)^2] = \sigma_x^2.$$

For a discrete r.v. with pf $f_x(x)$ we have

$$\sigma_x^2 = \sum_x (x - \mu_x)^2 f_x(x).$$

Fact: $\sigma^2 = E[(X - \mu_x)^2]$
 $= E(X^2) - [E(X)]^2$
 $= 2^{nd} \text{ moment} - 1^{st} \text{ moment squared}$
 $= \sum_x x^2 f(x) - (\sum_x x f(x))^2$ when X is discrete with pf $f(x)$.

Proof:

$$\begin{aligned}\sigma^2 &= \sum_x (x - \mu_x)^2 f(x) \\ &= \sum_x (x^2 - 2\mu_x x + (\mu_x)^2) f(x) \\ &= \sum_x x^2 f(x) - 2\mu_x \underbrace{\left(\sum_x x f(x)\right)}_{\mu_x} + \mu_x^2 \underbrace{\left(\sum_x f(x)\right)}_{=1} \\ &= E(X^2) - 2\mu_x^2 + \mu_x^2 \\ &= E(X^2) - \mu_x^2 = E(X^2) - (E(X))^2\end{aligned}$$

Variance measures the spread (or dispersion) in the histogram for a random variable. The measure of spread then, back in original units of x , is

Standard deviation, denoted by σ_x ,

and is

$$\sigma_x = + \sqrt{\sigma_x^2} = + \sqrt{\text{variance}(x)}.$$

Back to example on p. 3; we toss a "fair" coin three times. $X = \# \text{ of heads}$.

$$\begin{aligned} \mu_x = E(X) &= \sum_x x f(x) = 0\left(\frac{1}{8}\right) + 1\left(\frac{3}{8}\right) + 2\left(\frac{3}{8}\right) + 3\left(\frac{1}{8}\right) \\ &= \frac{3}{8} + \frac{6}{8} + \frac{3}{8} = \frac{12}{8} = \frac{3}{2} = 1.5 \end{aligned}$$

$$\begin{aligned} \mu_2 = E(X^2) &= \sum x^2 f(x) = 1^2\left(\frac{3}{8}\right) + 2^2\left(\frac{3}{8}\right) + 3^2\left(\frac{1}{8}\right) \\ \text{2nd moment} &= \frac{3}{8} + \frac{12}{8} + \frac{9}{8} \\ &= \frac{24}{8} = 3 \end{aligned}$$

$$\begin{aligned} \sigma_x^2 = \text{Var}(X) &= E(X^2) - (\mu_x)^2 \\ &= 3 - \left(\frac{3}{2}\right)^2 = 3 - \frac{9}{4} = 3 - 2.25 \\ &= .75 \end{aligned}$$

$$\text{Stand. dev.}(X) = \sigma_x = \sqrt{.75} \approx .866$$

Experiment

Keep rolling a "fair" die until a 6 occurs.
Let $X = \#$ of rolls (trials) needed to observe a 6.

~~$S = \{6\}$~~ Let $S = \text{success} = \text{a } 6 \text{ results}$
 $F = \text{failure} = 1, 2, 3, 4, \text{ or } 5 \text{ results.}$

We think of this as a sequence of identical Bernoulli Trials. A Bernoulli trial is one in which results in exactly one of two possible outcomes, S or failure. We can ~~assign~~ define a r.v. X as

$$X = \begin{cases} 1 & \text{if } S \text{ occurs} \\ 0 & \text{if } F \text{ occurs.} \end{cases}$$

Let $p = \text{prob. of success}$. Let $q = 1 - p = \text{prob. of failure}$.
Then X is called a Bernoulli r.v. with pf.

$$f_x(x) = p^x q^{1-x}, \quad x = 0, 1.$$

$$\left(\begin{array}{l} f_x(0) = P(X=0) = p^0 q^1 = q \\ f_x(1) = P(X=1) = p \end{array} \right),$$

$$\Omega = \{ S, FS, FFS, FFFS, \dots \}$$

So, $X = 1, 2, 3, \dots$

X is called a geometric r.v. as it counts the # of Bernoulli trials needed to obtain the first success. It's pf is

$$f_x(x) = P(X=x) = q \cdot p^{x-1}, \quad x=1, 2, 3, \dots$$

For the die problem, $p = P(\text{a 6 occurs}) = P(S) = \frac{1}{6}$

$$q = P(F) = \frac{5}{6}$$

$$f_x(x) = \left(\frac{5}{6}\right)^{x-1} \cdot \frac{1}{6}, \quad x=1, 2, 3, \dots$$

Note: For a geometric pf.

$$f_x(x+1) = q^{(x+1)-1} p = q^x p = q q^{x-1} p = q f_x(x)$$

i.e. $f(x+1) = q f(x)$

~~$$E(X) = \sum_{x=1}^{\infty} x f(x) = \frac{1}{q} \sum_{x=1}^{\infty} x f(x+1) = \frac{1}{q} \sum_{x=1}^{\infty} x q f(x)$$~~

$$E(X) = \frac{1}{p} \quad \text{and} \quad \text{Var}(X) = \frac{q}{p^2}$$